# A GENDER COMPARISON OF ECONOMISTS’ PUBLICATIONS 

E. Bruce Hutchinson, Marc A. Loizeaux, Leila J. Pratt, and Stephanie Smullen ${ }^{1}$


#### Abstract

An ordered probit model is used to examine the impact of gender and the quality of the PhD granting institution on the publication record of male and female economists who received $t$ heir doctorate in 1985. This analysis indicates th at men an d women have different publication patterns regardless of where they received their PhD and that the quality of the PhD granting institution has n om easurable effect on an individual's publication record.


Key Words: Gender, ordered probit model, journal publications
JEL Classification: J16

## Introduction

Numerous s tudies ( Davis, H uston a nd Patterson [ 2001], G oodwin a nd S auer [1995], H utchinson a nd Z ivney [ 1995], a nd Laband a nd P iette [ 1994]) e xamine $t$ he journal-publication be havior of i ndividuals $w$ ith a doctorate in e conomics. A subcategory of this literature is a continuing professional interest in the comparative journalpublication records of male and female economists. B ased on a 1966 National Science Foundation survey, H ansen, W eisbrod a nd S trauss report that w omen e conomists ha ve "higher av erage job quality" yet "lower research productivity (1978, p.737)." Fish and Gibbons whose research focused on journal publication between 1969 and 1986 conclude "...that men significantly out-publish w omen ...w hether the s amples [are] regarded as matched pairs or as two independent samples (p. 97)." McDowell and Smith, using data from 1968 t o 1975 f or an equal num ber of m ale a nd female e conomists from top 20 institutions, c onclude that on a verage, w omen produced fewer publ ications e ven after adjusting for the number of coauthors (1992, p. 75). Ginther and Kahn, relying upon National Science Foundation data for doctorates earned from 1974 through 2000, w rite that "notably, men publish more than women, particularly in non-top-10 journals (2004, p. 199)."

In a 2006 a rticle, McDowell , S ingell a nd S tater, using data from the A merican Economics Association (AEA) Directories for the years 1964, 1974, 1985, 1989, 1993, and 1997, conclude that by 1993 the conclusion of earlier studies that male economists

[^0]were $m$ ore likely to publish $m$ ore than their female counterparts (p. 166-67) was no longer applicable. Their analysis of post-1993 data indicates that males and females had substantively equal publication records.

The present study uses statistical analysis to compare male and female publication records a djusted b y the quality of t he PhD granting i nstitution. Quality here is determined by school tier as established by the National Research Council. We, as other authors excepting McDowell, Singell and Stater, find a continuing statistical difference in journal publication records when we simultaneously consider gender and the quality of the PhD granting institution.

## Data and Results

Our data are drawn from the 1985 and 1986 listing of "Doctoral Dissertations in Political Economy in American Universities and C olleges" p ublished in the D ecember 1985 a nd 1986 e ditions of $t$ he American Economic Review. Th ese lists id entify individuals and the year in which the PhD is conferred. Our data set includes all listed individuals who received a PhD in $1985 .^{2}$ The Economic Literature Database (Heck, 2001), which contains 250 -plus j ournals, was u sed to i dentify economics a nd related journal articles (hereafter "journal article") published by these individuals between 1985 and 1999. Counted were articles and notes; omitted were comments, replies, discussions, and book reviews, which is the general treatment followed in the literature. If articles were co-authored, even if both authors were from the 1985 PhD class, each was given credit for one publication. The gender of an individual was determined based on name and where necessary and possible by contacting the individual. ${ }^{3}$

The original data set contained 720 i ndividuals; however, we were unable to determine $t$ he gender of 50 individuals. Thus our $w$ orking da ta s et c ontains 670 individuals. Of these 115 or $17.2 \%$ were females and 555 were males. Three-hundred and twenty-seven ( $48.8 \%$ ) of these individuals published at least one journal article.

A larger percentage ( see Table 1) of women ( $58 \%$ ) failed to publish at least one article be tween 1985 and 1999 t han di dmen (50\%). However a slightly 1 arger percentage of women (13\%) than men (12\%) published exactly one article. Likewise the percentage of $w$ omen ( $8 \%$ ) w ho publ ished $t$ hree a rticles du ring $t$ his $p$ eriod $w$ as a lso somewhat 1 arger $t$ han $t$ he $p$ ercentage of $m$ en (5\%). However a s ubstantially $l$ arger percentage of $m$ en ( $7 \%$ ) than women ( $3 \%$ ) publ ished exactly two a rticles and an even higher pe rcentage of $m$ en ( $26 \%$ ) t han $w$ omen ( $18 \%$ ) publ ished four or $m$ ore a rticles during this period.
${ }^{2}$ The lis ts in clude individuals who e arn t he P hD from C anadian U niversities. T hese individuals are omitted from our sample because their PhD is from a non-U.S. university and the tier rankings used include only U.S. universities.
${ }^{3}$ Many faculty and students with knowledge of foreign languages and cultures assisted in this $d$ etermination. $E$-mails w ere a lso sent in a $n$ e ffort to $d$ etermine the $g$ ender of individual economists.

| Table 1 |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Number of Publications by Gender |  |  |  |  |  |
| Publications | 0 | 1 | 2 | 3 | $4+$ |
| Female | 67 | 15 | 3 | 9 | 21 |
|  | $58 \%$ | $13 \%$ | $3 \%$ | $8 \%$ | $18 \%$ |
| Males | 276 | 66 | 40 | 29 | 144 |
|  | $50 \%$ | $12 \%$ | $7 \%$ | $5 \%$ | $26 \%$ |
| Total | 343 | 81 | 43 | 38 | 165 |
|  | $51 \%$ | $12 \%$ | $6 \%$ | $6 \%$ | $25 \%$ |

Universities were p laced into school tiers a ccording to the 1982 r anking of economics departments as reported inthe a ppendices of $t$ he 1995 N ational R esearch Council update to the 1982 assessment of research-doctorate programs. Essentially, we used Hansen's first (highest or be st) to f ifth (lowest) tier designations to $g$ roup the schools. ${ }^{4}$ Table 2 shows this stratification.

As can be s een, $40 \%$ of our sample received their PhD from a tier 1 or tier 2 institutions while $32 \%$ r eceived $t$ heir de gree from a $t$ ier 5 s chool. $14 \%$ of $t$ he $m$ ales received $t$ heir de grees from a t ier 1 s chool compared to onl y $8 \%$ of the females. In addition, a hi gher $p$ ercentage of $w$ omen $t$ han $m$ en $g$ raduated $f$ rom $t$ ier 4 or $t$ ier 5 institutions.

[^1]Tier 5: All other Colleges and Universities.

| Table 2 |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | Tier 1 | Tier 2 | Tier 3 | Tier 4 | Tier 5 |
| Publications | 1 | 2 | 3 | 4 | 5 |
| Females | 9 | 28 | 16 | 21 | 41 |
|  | $8 \%$ | $24 \%$ | $14 \%$ | $18 \%$ | $36 \%$ |
| Males | 76 | 156 | 90 | 62 | 171 |
|  | $14 \%$ | $28 \%$ | $16 \%$ | $11 \%$ | $31 \%$ |
| Total | 85 | 184 | 106 | 83 | 212 |
|  | $13 \%$ | $27 \%$ | $16 \%$ | $12 \%$ | $32 \%$ |

Table 3 s hows publishing records s tratified b y gender a nd s chool tier. Giving consideration $t$ ot he $s$ ensitivity of pe rcentages ba sed on $s$ mall num bers, no pa ttern differences ar e r eadily observed. An or dered probit m odel is us ed to determine the impact of these variables on the probability that an individual will publish $1,2,3$ or 4 or more articles between 1985 and 1999. In general, this model takes the form:

$$
\mathrm{y}^{*}=\boldsymbol{\beta}, \mathbf{x}+\varepsilon .
$$

$\mathrm{y}^{*}$ is not observed but we do observe

$$
\begin{array}{ll}
\mathrm{y}=0 & \text { if } \mathrm{y}^{*} \leq 0 \\
\mathrm{y}=1 & \text { if } 0 \leq \mathrm{y}^{*} \leq \mu_{1} \\
\mathrm{y}=2 & \text { if } \mu_{1} \leq \mathrm{y}^{*} \leq \mu_{2} \\
& \cdot \\
\mathrm{y}=\mathrm{j} & \cdot \\
& \\
& \text { if } \mu_{\mathrm{j}-1} \leq \mathrm{y}^{*}
\end{array}
$$

The $\mu$ 's are unknown parameters and are estimated with the $\boldsymbol{\beta}$ 's. The values of both these parameters depend on the set of measurable factors, $\mathbf{x}$, and the unobservable factors $\varepsilon$. The error term, $\varepsilon$, is assumed to be normally distributed across observations. It is standardized to a mean of zero and a variance of one. The resulting normal distribution gives us the following probabilities:

$$
\begin{aligned}
& \operatorname{Prob}(\mathrm{y}=0)=\Phi(-\boldsymbol{\beta} \mathbf{x}), \\
& \operatorname{Prob}(\mathrm{y}=1)=\Phi\left(\mu_{1}-\boldsymbol{\beta}, \mathbf{x}\right)-\Phi(-\boldsymbol{\beta}, \mathbf{x}), \\
& \operatorname{Prob}(\mathrm{y}=2)=\Phi\left(\mu_{2}-\boldsymbol{\beta}, \mathbf{x}\right)-\Phi\left(\mu_{1}-\boldsymbol{\beta} \mathbf{x}\right), \\
& \operatorname{Prob}(\mathrm{y}=3)=\Phi\left(\mu_{3}-\boldsymbol{\beta}, \mathbf{x}\right)-\Phi\left(\mu_{2}-\boldsymbol{\beta} \mathbf{x}\right), \\
& \operatorname{Prob}(\mathrm{y}=4)=1-\Phi\left(\mu_{3}-\boldsymbol{\beta}, \mathbf{x}\right) .
\end{aligned}
$$

For all the probabilities to be positive:

$$
0<\mu_{1}<\mu_{2}<\mu_{3}
$$

| Table 3 |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Number of Publications by School Tier and Gender |  |  |  |  |  |
| Publications | 0 | 1 | 2 | 3 | $4+$ |
| Tier 1 |  |  |  |  |  |
| Females | 4 | 2 | 1 | 1 | 1 |
|  | $44 \%$ | $22 \%$ | $11 \%$ | $11 \%$ | $11 \%$ |
| Males | 34 | 10 | 10 | 3 | 19 |
|  | $45 \%$ | $13 \%$ | $13 \%$ | $4 \%$ | $25 \%$ |
| Tier 2 |  |  |  |  |  |
| Females | 18 | 3 | 0 | 1 | 6 |
|  | $64 \%$ | $11 \%$ | $0 \%$ | $4 \%$ | $21 \%$ |
| Males | 85 | 13 | 13 | 10 | 35 |
|  | $54 \%$ | $8 \%$ | $8 \%$ | $6 \%$ | $22 \%$ |
| Tier 3 |  |  |  |  |  |
| Females | 6 | 2 | 1 | 1 | 6 |
|  | $38 \%$ | $13 \%$ | $6 \%$ | $6 \%$ | $38 \%$ |
| Males | 47 | 11 | 6 | 5 | 21 |
|  | $52 \%$ | $12 \%$ | $7 \%$ | $6 \%$ | $23 \%$ |
| Tier 4 |  |  |  |  |  |
| Females | 12 | 0 | 0 | 4 | 5 |
|  | $57 \%$ | $0 \%$ | $0 \%$ | $19 \%$ | $24 \%$ |
| Males | 31 | 8 | 3 | 5 | 6 |
|  | $58 \%$ | $15 \%$ | $6 \%$ | $9 \%$ | $11 \%$ |
| Tier 5 |  |  |  |  |  |
| Females | 27 | 8 | 1 | 2 | 3 |
|  | $66 \%$ | $20 \%$ | $2 \%$ | $5 \%$ | $7 \%$ |
| Males | 79 | 24 | 8 | 6 | 54 |
|  | $46 \%$ | $14 \%$ | $5 \%$ | $4 \%$ | $32 \%$ |

The e xplanatory va riables or t he regressors, $\mathbf{x}$, a re G ENDER which e quals 1 i ft he individual is male, a nd a s et of dum my $v$ ariables that de signate $t$ he $t$ ier of $t$ he P HD granting institution. Table 4 reports the results of the ordered probit. None of the school tier dummies are significant indicating that the quality of the PhD granting institution is unimportant in explaining an individual's probability of publishing. However, gender is positive and significant indicating that males have a significantly hi gher probability of publishing than females.

| Table 4 <br> Ordered Probit Regression Results |  |  |  |
| :---: | :---: | :---: | :---: |
|  |  |  |  |
| Variable | Coefficient | Error | t-Statistic |
| Constant | -0.1858 | 0.1277 | -1.455 |
| Gender | 0.2313 | 0.1211 | 1.910 |
| Tier 1 | 0.0310 | 0.1465 | 0.212 |
| Tier 2 | -0.1349 | 0.1175 | -1.149 |
| Tier 3 | -0.0106 | 0.1371 | -0.077 |
| Tier 4 | -0.1367 | 0.1498 | -0.913 |
| $\mu_{1}$ | 0.3112 | 0.0326 | 9.546 ** |
| $\mu_{2}$ | 0.4882 | 0.0402 | 12.144 ** |
| $\mu_{3}$ | 0.6592 | 0.0464 | 14.195 ** |
| * | Significant at the $5 \%$ level Significant at the $10 \%$ level |  |  |
| ** |  |  |  |

As usual in models with discrete dependent variables the marginal impacts of the $\mathbf{x}$ values on the probabilities are not equal to the coefficients. To measure the marginal impact of a binary explanatory variable, one must compare the probabilities that result when the variable takes on its two values ( 0 or 1 ) with all other variables held constant at their means. Table 5 shows the marginal impact of each of the binary regressors (school tiers or gender) on the five publishing probabilities given all the other variables are held constant at $t$ heir $s$ ample $m$ eans. For e xample, a ccording to the estimated ma rginal impacts presented in Table 5, m ales are $9.20 \%$ more likely to not publish than females given the school tiers are held constant at their mean values. Likewise a graduate of a Tier 2 institution is $5.40 \%$ less likely to not publish than other individuals given gender and the other three tier dummy variables are held constant at their mean values.

The $s$ mall $d$ ifferences in ma le a nd female $p$ ublication $p$ robabilities in the $f$ ive school $t$ iers $r$ einforce $t$ he non -significance of $t$ he $t$ ier $d$ ummy $v$ ariables. $T$ he $g$ ender dummy va riables however pr ovide some s urprising results. M en ar e about $9 \%$ more likely than women to not publish at all. Women, on the other hand, are about $8 \%$ more likely than men to publ ish 4 or more articles or to be "super publishers". In a ddition, women and men are almost equally likely to publish 1, 2, or 3 articles between 1985 and 1999.

## Table 5

Marginal Impact of Gender \& School Tier on Publishing Probabilities

|  | Prob. $\mathrm{y}=0$ | Prob. $\mathrm{y}=1$ | Prob. $\mathrm{y}=2$ | Prob. $\mathrm{y}=3$ | Prob. $\mathrm{y}=4$ |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Gender $=0$ | 0.412 | 0.123 | 0.069 | 0.064 | 0.331 |
| Gender $=1$ | 0.504 | 0.122 | 0.065 | 0.058 | 0.252 |
| Difference | 0.092 | -0.001 | -0.004 | -0.006 | -0.079 |
|  |  |  |  |  |  |
| Tier 1 = 0 | 0.486 | 0.122 | 0.066 | 0.060 | 0.266 |
| Tier 1 = | 0.499 | 0.122 | 0.065 | 0.058 | 0.256 |
| Difference | 0.013 | 0.000 | -0.001 | -0.002 | -0.010 |
|  |  |  |  |  |  |
| Tier 2 = 0 | 0.503 | 0.122 | 0.065 | 0.058 | 0.253 |
| Tier 2 = | 0.449 | 0.124 | 0.068 | 0.062 | 0.298 |
| Difference | -0.054 | 0.002 | 0.003 | 0.004 | 0.045 |
|  |  |  |  |  |  |
| Tier 3 = 0 | 0.488 | 0.123 | 0.066 | 0.059 | 0.264 |
| Tier 3 = 1 | 0.484 | 0.123 | 0.066 | 0.059 | 0.264 |
| Difference | -0.004 | 0.000 | 0.000 | 0.000 | 0.000 |
|  |  |  |  |  |  |
| Tier 4 $=0$ | 0.488 | 0.123 | 0.066 | 0.059 | 0.264 |
| Tier 4 = 1 | 0.483 | 0.123 | 0.066 | 0.059 | 0.269 |
| Difference | -0.005 | 0.000 | 0.000 | 0.000 | 0.005 |

## Conclusion

This study used research records based on a search of the Economic Literature Database citations to 250 -plus economics and related journals to examine the impact of gender and school tier on $t$ he probability that an individual who received the PhD in 1985 w ould publish zero, one, two, three or four or more articles between 1985 and 1999. We find that for this group, the tier (general quality) of the graduate institution from which they graduated does not impact their publishing probability. W e also find that men have a significantly different publication pattern than women though the difference is narrow. In particularly we find that men in this cohort are more likely to not publish at all. On the other hand, we find that women are more likely to be "super publishers" publishing four or more articles between 1985 and 1999.

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[^0]:    ${ }^{1}$ E. Bruce Hutchinson is Professor of Economics, Department of Economics, University of Tennessee at Chattanooga, Chattanooga, TN 37403; Marc A. Loizeaux is Manager, Provider P erformance Assessment A nalytics, BlueCross B lueShield of T ennessee, Chattanooga, T N 37402 ; Leila J. P ratt is Hart Professor of E conomics, University of Tennessee at Chattanooga; a nd, Stephanie S mullen, is Professor of C omputer S cience, Computer Science, University of Tennessee at Chattanooga.

[^1]:    ${ }^{4}$ Schools in the various Tiers are:
    Tier 1: Chicago, Harvard, MIT, Princeton, Stanford, Yale
    Tier 2: Columbia, Michigan, M innesota, N orthwestern, Pennsylvania, R ochester, U CBerkeley, UCLA, UW-Madison
    Tier 3: Brown, C al-Tech, C arnegie-Mellon, C ornell, D uke, Illinois, J ohns H opkins, Maryland, Michigan State, New Y ork University, North Carolina, UC-San Diego, Virginia, Virginia Polytechnic Institute, Washington-Seattle
    Tier 4: Boston University, C laremont, Florida, Iowa, Iowa State, Massachusetts, Ohio State, Pennsylvania State, Pittsburgh, Purdue, SUNY-Stony Brook, T exas A \&M, Texas-Austin, UC-Davis, UC -Santa B arbara, U SC, V anderbilt, W ashington-St. Louis

